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Abstract

Using data from the British Household Panel Survey (BHPS) we show performance pay (PP) increased earnings dispersion among men and women, and to a lesser extent among full-time working women, in the decade of economic growth which ended with the recession of 2008. PP was also associated with some compression in the lower half of the wage distribution for women. The effects were predominantly associated with a broad measure of PP that included bonuses. However, these effects were modest and there is no indication that PP became increasingly prevalent, as some had predicted, over the decade prior to recession.

Key words: wages, wage inequality, performance pay, bonuses

JEL classification: J31; J33

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1. Introduction

Income inequality has grown in English-speaking economies in recent decades, largely due to growing wage inequality (Atkinson, Piketty and Saez, 2011). A variety of explanations have been proffered, including increasing returns to skill induced by skills-biased technological change (SBTC) (Autor, Katz and Kearney, 2008), changes in labour market institutions, most notably de-unionisation (Dustmann et al., 2009; Card, Lemieux and Riddell, 2004) and increased trade (Autor et al., 2013). In their seminal paper for the United States Lemieux, MacLeod and Parent (LMP) (2009) show that performance pay (PP) accounted for one-fifth of the growth in wage inequality among men between the late 1970s and early 1990s, and most of the growth in wage inequality among high earners in the top quintile. They show that PP became more widespread between the 1970s and early 1990s, was closely tied to individuals' productive characteristics, and that the returns to these characteristics were rising faster in PP jobs than in fixed wage jobs. Their findings are consistent with a world in which SBTC increases the rewards for more productive workers and induces firms to resort to PP to both attract and incentivise those workers.

LMP's (2009) model, which draws on the work of Lazear (1986; 2000) and Prendergast (1999), indicates PP generates higher wage dispersion than fixed rate pay (FP) due to the sorting of high ability workers into PP jobs - a labour market segmentation type argument - and because PP reflects individuals' marginal product more accurately than fixed wage schedules. Growth in PP jobs allows high ability workers to recoup returns to their ability in a way that is not possible with fixed wages, while the higher incidence of PP at the top end of the earnings distribution will also generate higher wage dispersion.

LMP attribute the increased use of PP to SBTC and the declining costs of worker monitoring due to advances in technology. These trends are likely to have continued in the period since the mid-1990s which LMP were studying, both in the United States and in other industrialised countries. For instance, Sommerfeld (2013) documents an almost continuous rise in the share of PP jobs between 1984 and 2009.

And yet LMP's findings have recently been challenged in a series of papers using data for the United States. Using establishment data from the Bureau of Labor Statistics' Employer Costs for Employee Compensation (ECEC) series (which derives from the National Compensation Survey) Gittleman and Pierce (2013) show the proportion of jobs with PP rose in the 1990s, only to fall in the 2000s such that, by 2013, it had *declined* by about one-fifth since LMP's study period, irrespective of how one measures PP. This decline is apparent throughout the wage distribution but is concentrated among low earners. Furthermore, in a second paper, Gittleman and Pierce (2015) show the contribution of PP to growth in the earnings distribution in the first decade of the 21st Century has been small - in the order of 9 per cent of the growth in variance. Sommerfeld's analysis for Germany also showed that despite the expansion of PP, it did not lead to increased wage inequality because it was associated with higher wages across the board and not just for high earners.

Two more papers find LMP's basic results do not hold for some parts of the working population. Like LMP, Heywood and Parent (2012) analyse the Panel Survey of Income Dynamics (PSID). They find that, during the period 1976-1998, the tendency for PP to be associated with greater wage inequality at the top of the male earnings distribution applies to white workers but not to black workers. In a second paper using the National Longitudinal

Survey of Youth (NLSY), Heywood and Parent (2013) find skilled fathers select into PP jobs, whereas skilled mothers select out of PP jobs, a finding which is not consistent with standard assumptions regarding workers sorting into PP jobs on ability. This, in turn, raises questions about the effects of PP on wage inequality.

In Britain wage inequality among full-time workers has been rising since the late 1970s, although the rate of change slowed dramatically in the 2000s, with all the growth being confined to the top part of the wage distribution (Machin, 2011; Lindley and Machin, 2013). Over the whole period the graduate wage premium rose, despite growth in the graduate share in employment and hours, suggesting demand for highly skilled labour was exceeding its supply (Lindley and Machin, *op. cit.*). This is consistent with SBTC, and the authors find direct evidence of greater demand for more educated workers in more technologically advanced industries (*op. cit.*: 175-176). They also point to the introduction of the national minimum wage in 1999 and its subsequent up-rating as a reason for the stability in the 50-10 wage differential in the 2000s.

Although they point to the potential importance of SBTC in the British context, Lindley and Machin do not consider the potential role played by PP in growing wage inequality. There is some evidence that annual bonuses have contributed to an increase in wage inequality at the top of the earnings distribution in the last decade or so, primarily as a result of large bonus receipt by bankers, traders and other well-paid professionals in the Finance sector (Bell and Van Reenen, 2010, 2013).¹ These employees may be sharing in the substantial rents

¹ However, bonuses account for only a small proportion of total earnings for those outside the top decile of earners (Bell and Van Reenen: 2013, 10-11).

generated by a lack of competition in the sector. Alternatively, they may be benefiting from productivity "scaling" effects that accrue to highly productive employees when changes such as increased firm size and capital intensification "scale up" worker productivity, increasing returns to their employer. This is the type of effect identified by Gabaix and Landier (2008) and Kaplan (2012) in relation to "superstars" such as CEOs.

But, aside from the effects of bonus payments at the very top, what effects has PP had on the overall wage distribution in Britain? Two studies using cross-sectional linked employer-employee data come to different conclusions. Manning and Saidi (2011) show that, although there is a wage premium attached to the receipt of PP, it had a negligible effect on wage dispersion in 2004. However, using data from the 2011 Workplace Employment Relations Survey Bryson et al. (2014) find PP results in a sizeable widening in wage differentials relative to a counterfactual wage distribution, and that this effect is larger higher up the earnings distribution. The premium rises markedly as one moves up the hourly wage distribution: it is seven times higher at the 90th percentile than it is at the 10th percentile in the wage distribution (.42 log points compared to .06 log points). This, coupled with the higher incidence of PP among those with wage-enhancing attributes, means PP contributes substantially to higher wage dispersion in Britain. However its overall effect on the wage distribution is less marked than it might have been due to the relatively low proportion of employees on PP contracts in Britain.

This paper contributes to the literature in a number of ways. First, in light of the debate in the US about the changing role of PP, we track the incidence of PP using the British Household Panel Survey (BHPS) for the period 1998-2008 that immediately preceded the

recession. We consider alternative broad and narrow definitions of PP and estimate their individual, job and workplace correlates. Next we estimate the premium associated with PP jobs and look at how it changed over the decade. Finally we estimate the effects of PP on wage dispersion and changes in the wage distribution over the period, accounting for the changing PP premium as well as changes in the prevalence of PP at different parts of the distribution.

We find no indication that PP jobs (broadly or narrowly defined) are becoming increasingly prevalent. Depending on the measure used and splitting by gender, we find either gradual declines or broad stability, although PP jobs may have picked up slightly among full-time women in the two years before the recession. Nonetheless the returns to PP remain positive, even when controlling for unobserved personal characteristics, and in fact seem to have increased over the period. Confirming other recent studies, we show that wage inequality grew somewhat during the decade of economic growth that ended abruptly with the recession, largely due to growing earnings dispersion in the top half of the wage distribution (with some reduction in inequality at the bottom for women). Estimates of PP effects on the counterfactual wage distribution confirm PP increased earnings dispersion among men and women, including the sub-group of full-time working women. PP also appears to have contributed to reduced wage dispersion at the bottom among women. In both cases, the changes are largest for the broad measure of PP, which includes bonuses. Nevertheless the effects overall are reasonably modest – while overall PP remains a disequalising force on the wage distribution in Britain, the fact that it has not become more widespread has limited its impact on wage inequality.

In the next section we outline the theoretical links between wage dispersion and PP. Section Three then introduces the data, followed by section Four which presents results relating to the incidence and correlates of PP followed by its links to wages and wage dispersion in Britain. Finally section Five discusses the implications of the findings and draws some conclusions.

2. Wage Dispersion and Performance Pay

In perfectly competitive labour markets in which firms and workers have perfect information employees would be paid their marginal product, that is, they would be paid for their performance. However, employers and employees often prefer fixed wage contracts based on time rather than effort or output. Employers may find fixed wages less costly to administer, especially if labour inputs or outputs are costly to monitor: it can be costly for firms to identify the contribution of individual employees to output, while factors beyond the control of the employee, and even the firm, mean output is affected by factors other than employees' talent and effort. In standard economic theory wage dispersion rises when employees are paid for their performance, compared to a counterfactual scenario in which they are paid a fixed wage. Under fixed wage schedules employees are paid for time worked, whereas under PP they are paid for output. Heterogeneity in individuals' ability to increase output, either by virtue of talent or effort, is ignored in fixed wage schedules, but it does have a bearing on earnings when pay is linked to performance.

There are three channels that may lead to higher earnings dispersion in the presence of PP. The first is a mechanical effect: PP reveals underlying differences in individuals' productivity that were previously ignored. Second, PP may have the effect of incentivising effort:

employees can raise (lower) their earnings through higher (lower) effort such that variance in effort induces variance in earnings, whereas employees' earnings are not a function of effort in fixed wage jobs. Third, employees will sort into (out of) PP jobs according to talent and other traits (such as their tastes for effort and risk) that may affect their earnings. If more able workers sort into PP jobs where they can command higher earnings, while less able workers prefer the guarantee of a fixed wage, the market will segment into high and low earners along PP lines. Thus via all three of these channels, the introduction of PP should lead to greater wage dispersion than might obtain if all workers were paid a fixed wage.²

Of course this is an over-simplistic picture because job retention and job progression are often performance-related, even when workers are paid a fixed wage, because wage levels and earnings progression reflect workers' efforts and talent, while career concerns can incentivise effort (Prendergast, 1999; Papps et al., 2011). But the link between performance and pay is usually more explicit and more direct in the presence of PP schemes.

While all these considerations suggest that PP will be associated with greater wage dispersion in cross section, the impact of PP on changes in wage dispersion are less clear. LMP (2009: 3-4) discuss some reasons why PP may induce growth in earnings dispersion. If demands for more skilled and more able workers are rising due to SBTC or globalisation, this will raise the market value of more talented workers such that firms may bid up their price relative to less talented workers as they try to influence the job matching process. This, in turn, may induce greater worker sorting between PP and FP jobs, contributing to growth in the

² If there is strong sorting on ability into (out of) the PP sector, there may be greater homogeneity among the workers in the PP sector than in the fixed wage sector which may result in wage dispersion within the PP sector being lower than dispersion in the fixed wage sector.

dispersion of earnings between PP and fixed pay jobs. If there is an increase in the prevalence of PP, particularly at the top end of the earnings distribution, this will also contribute to a growth in earnings dispersion.

3. Data

We analyse data from the British Household Panel Survey (BHPS), which began in 1991 with a sample of some 5,000 households from England, Scotland and Wales. Household members aged 16 or over were interviewed annually (usually in September or October) through to the final year 2008. The sample consists of the original “Essex sample” of households together with associated new entrants to the survey (e.g. children reaching age 16 and new partners of original sample members), thus keeping it broadly representative of the British population structure.³

The analyses focus on employees aged 18-64 years, excluding those reporting total weekly hours of 100 or more or 5 or less. This restriction, which affects fewer than 2% of observations, reduces possible measurement error in hourly wages arising from extreme reports of hours worked. It also eliminates very small jobs. To account for the possibility of different wage determination processes across gender, we perform separate analyses for men and women. We also analyse a sample of women in full-time jobs only to make sure that any gender differences do not reflect the much larger proportion of women in part-time jobs (where PP is less common). All our estimates are weighted using the cross-sectional weights

³ Given the rules for new entrants to the survey, it will not reflect the impact of immigration in the mid-2000s. In addition, we do not include the Scotland, Wales and Northern Ireland extension samples added in 1999 and 2001, as this would introduce a discontinuity in the data.

provided with the survey, which account for survey design and the likelihood that a respondent appears in a particular wave.

As is standard in the literature our wages measure is hourly wages, which we compute as $(\text{usual gross pay}/(\text{usual basic hours}+1.5 \times \text{usual paid overtime}))$. The usual gross pay variable includes regular bonuses, commission and tips, so the hourly wage measure will take account of these components of PP. We can also construct a second hourly pay measure including more irregular bonuses (such as seasonal bonuses), derived from a separate question in the survey. While in principal this second measure better reflects total bonus payments received, it carries a risk of double counting if respondents report some bonus payments in answer to both questions (the second question does not explicitly exclude all regular bonuses). As a result we use the first wage measure as our baseline dependent variable, but as a robustness check we also run all analyses with the wage measure including irregular bonuses (the results are almost identical).

BHPS contains two measures of PP. The first, relating to bonuses, is derived from the question: **"In the last 12 months have you received any bonuses such as a Christmas or quarterly bonus, profit-related pay or profit-sharing bonus, or an occasional commission?"**. The second measure relates specifically to performance-related pay (PRP). Respondents are asked: **"Does your pay include performance related pay?"** The bonus question was asked in Waves 6-18 and the PRP question in Waves 8-18. As we wish to

combine information from the two measures we focus on Waves 8-18 covering the period 1998-2008.⁴

Gittleman and Pierce (2013) emphasise the importance of recognising that PP measures often capture different types of PP, some more closely related to individual productivity than others. In our data, the PRP question arguably captures pay linked to individual performance, while the bonus question captures payments like Christmas bonuses and rewards, such as profit related pay, that are probably linked to team or firm performance. Across the pooled sample PRP is roughly half as prevalent as bonus receipt (15% compared to 32%, see Table 1).

Table 1 – Bonus and PRP receipt (row and column percentages)

		PRP		
		No	Yes	Total
Bonus	No	90.8	9.2	100.0
		<i>73.2</i>	<i>41.2</i>	<i>68.3</i>
	Yes	71.7	28.3	100.0
		<i>26.8</i>	<i>58.8</i>	<i>31.7</i>
Total		84.8	15.3	100.0
		<i>100.0</i>	<i>100.0</i>	<i>100.0</i>

While the two measures are positively correlated, over 70% of those receiving bonuses do not get PRP and 41% of those receiving PRP do not get bonuses (Table 2). So to some extent PRP and bonus receipt are distinct types of compensation. As we show later it is also the case that employees in PRP and bonus jobs differ somewhat in their characteristics.

⁴ A different question about bonuses was asked in Waves 1-5: "Does your pay ever include incentive bonuses or profit related pay?" It includes fewer bonus components than the later question and about 10% fewer people reported bonuses each year by this measure. As we wish to use consistent measures for each

Table 2 – Overlap of bonus and PRP receipt (cell percentages)

		PRP		
		No	Yes	Total
Bonus	No	62.0	6.3	68.3
	Yes	22.7	9.0	31.7
	Total	84.8	15.3	100.0

Gittleman and Pierce (2012) present two PP measures: a broad measure including incentive pay and *all* bonuses; and a narrow PP measure restricted to incentive pay and *performance-based* bonuses only. We follow their approach. Our broad measure of PP combines the PRP and bonus questions, while our narrow measure is confined to the PRP measure. Table 2 shows that 62% of employees did not get either PRP or bonus, thus the prevalence of PP broadly measured is 38%, whereas 15% receive PP narrowly defined.

Tables 1 and 2 above relate to the *receipt* of PP. However, throughout the analysis presented in Section Four we follow other papers in the literature by focusing on PP *jobs*, not receipt. A job is a period of employment with the same employer in the same “grade”, i.e. if they get promoted it is a new job. A job is classified as a PP job where the respondent has been in receipt of PP on at least one occasion. This adjustment is made in recognition of that fact that some jobs are PP jobs but that, for whatever reason (poor performance on the part of the firm or individual, for instance) there has been no receipt of PP in a particular year - that is to say, the respondent may be in a job that pays for performance but, in a given year, the PP due is £0, thus making it hard to distinguish from a fixed pay job.

year, and also combine the bonus and PRP measures for some of the analysis, we do not use the earlier bonus question.

Since the probability of a PP job paying out for performance is partly a function of the number of times that job is observed in the data (which is lower for jobs near the ends of the data window) it is necessary to make an “endpoint adjustment” which accounts for the presence of jobs of different durations. Following LMP’s approach we construct an adjusted measure of the prevalence of PP jobs by estimating probit models for the probability of appearing in each wave of the data based on the number of times a job is observed. The resulting predicted probabilities are used to construct a weight which then effectively holds the distribution of the number of times a job is observed to that observed in the middle of our sample.

4. Results

4.1 The Prevalence of PP and Its Correlates

Figure 1a: Incidence of PP Jobs, Broad Measure

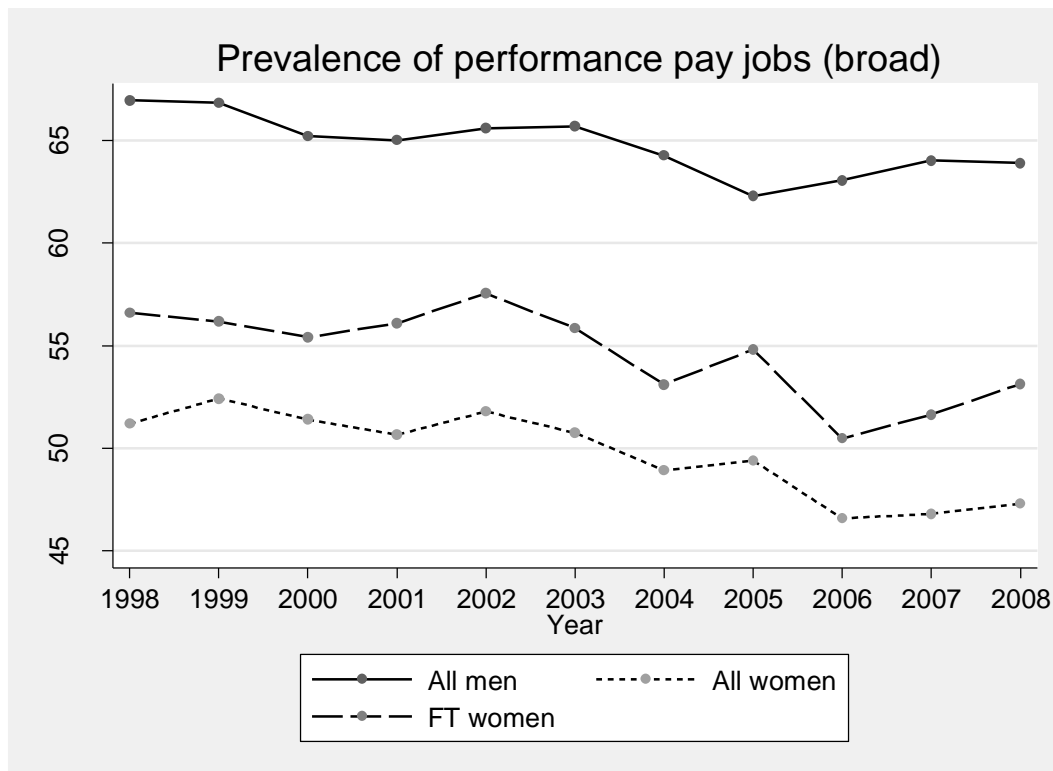


Figure 1a shows the incidence of PP jobs using the broad measure (PRP plus bonuses). It is apparent that men are more likely to be in PP jobs than women and that, among women, PP jobs are more common in full-time jobs. Throughout the period a little under two-thirds of jobs undertaken by men were PP jobs. Among women, the figure is closer to one-half. The incidence of PP jobs declines a little over the decade before the recession began, the drop being particularly notable among full-time women. There is no evidence at all that PP jobs became more common except perhaps a small rebound among women towards the end of the period.

Figure 1b: Incidence of PP Jobs, Narrow Measure

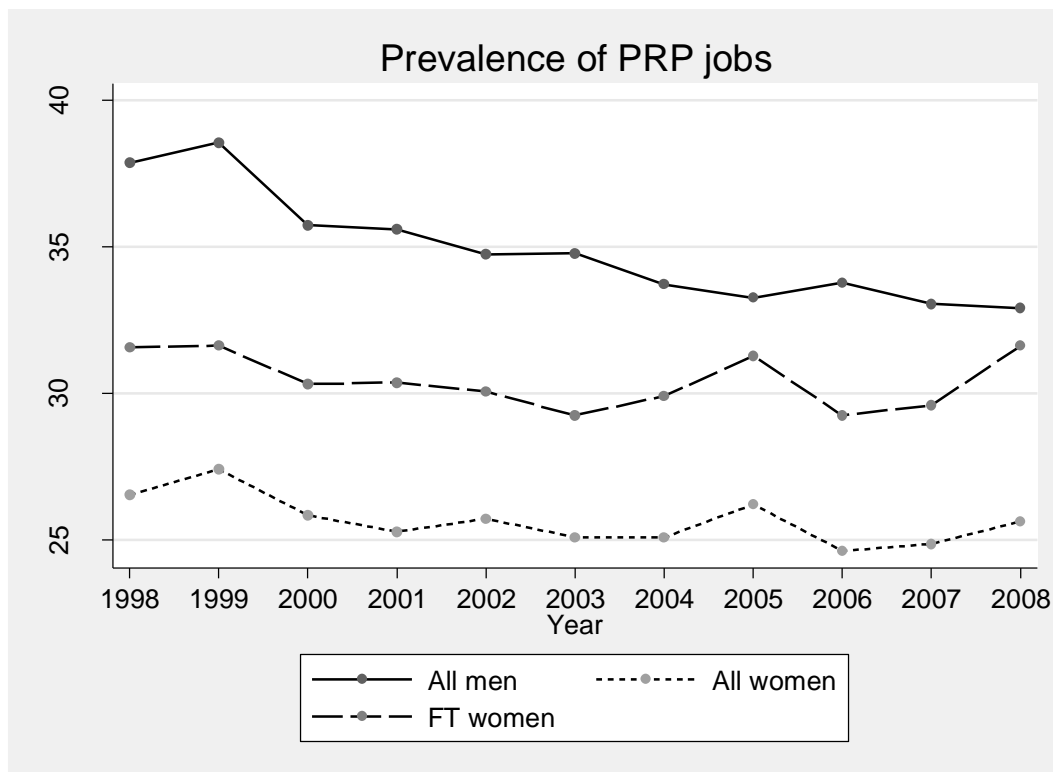


Figure 1b presents the same information but this time for the narrow measure of PP which is based solely on the performance-related pay question referred to in Section Three. The

male-female gap is smaller on this narrow measure, and it narrows over the period since 1998 because the percentage of jobs undertaken by men that are PP jobs has been falling. The percentage of PP jobs among women was broadly stable, though as for the broad measure, there is some indication that it may have increased slightly after 2006. Overall, though, we see no clear evidence that PP jobs expanded over this period.

Table 3 shows a bivariate probit estimated for the pooled sample which establishes the correlates of bonus-paying jobs and PRP jobs, having accounted for the positive and statistically significant correlation in unobservables between the two. There are a number of points worth noting. First, consistent with the graphical evidence, the incidence of both bonus jobs and PRP jobs has declined significantly since the turn of the century having conditioned on employees' demographic, job and workplace characteristics. For PRP jobs there was an abrupt decline in 2000 followed by further decline after 2006. Bonus jobs also fell sharply in 2000 but then declined more steadily. Second, those in receipt of both types of PP have higher gross hourly earnings than those in fixed pay jobs: even after accounting for occupation, tenure, and other characteristics that influence wages (such as firm size) those in PP jobs have gross hourly wages that are around one-third higher than those among observationally equivalent fixed pay employees.

Third, the male-female differential in PP jobs, apparent in the figures above, is not significant having accounted for other factors. PP jobs are more likely to be full-time, permanent, and in managerial, clerical and sales occupations. PRP jobs are more likely to be unionised than fixed wage jobs, but this is not the case for bonus jobs. The quadratic in years of job tenure turns at about 18 years for bonus jobs and 15 years for PRP jobs, both of

which are above the 90th percentile of the job tenure distribution, so the probability of bonuses increases in tenure for most employees. Unsurprisingly both types of PP job are more likely in larger organizations and the industry patterns are as found in the literature.

Table 3 – Correlates of bonus and PRP jobs

	Bonus job	PRP job	Wald test of equal effects
Female	-0.031	0.022	
Age/10	0.208**	0.106	
(Age/10) ²	-0.029**	-0.023*	
Married	-0.028	0.016	
Log (gross hourly wage)	0.331**	0.359**	
Part-time (<=30 hours total)	-0.145**	-0.266**	**
Seasonal or temporary job	-0.765**	-0.525**	**
Fixed term job	-0.624**	-0.596**	
Job tenure	0.109**	0.061**	**
Job tenure squared	-0.003**	-0.002**	**
TU or staff assoc at workplace	0.019	0.234**	**
Public sector	-0.507**	-0.064	**
Manager	0.139**	0.151*	
Professional	-0.069	0.152*	**
Technician	-0.030	0.025	
Clerical	0.111	0.070	
Craft	-0.051	0.049	
Personal	-0.195**	-0.14	
Sales	0.281**	0.425**	
Operative	-0.099	-0.057	
Routine	0	0	
Agriculture	-0.385**	-0.297*	
Mining & utilities	-0.03	0.329**	*
Construction	-0.215**	-0.190**	
Retail and hotels	-0.027	-0.055	
Communications	0.158**	0.041	
Finance and property	-0.001	0.168**	**
Other industries	-0.275**	-0.080	**
Social work & health	-0.851**	-0.741**	
Education	-0.891**	-0.243**	**
Public administration	-0.685**	-0.111	**
Manufacturing	0	0	
Establishment 500+ employees	0.152**	0.197**	

Establishment 50-499 employees	0.070**	0.099**	
Establishment 1-49 employees	0	0	
1999	-0.073**	0.02	**
2000	-0.156**	-0.067**	**
2001	-0.152**	-0.073**	**
2002	-0.168**	-0.089**	**
2003	-0.177**	-0.083**	**
2004	-0.226**	-0.106**	**
2005	-0.264**	-0.095**	**
2006	-0.309**	-0.130**	**
2007	-0.344**	-0.170**	**
2008	-0.369**	-0.192**	**
Constant	-0.761**	-1.618**	**
Correlation of equation errors	0.485**		
Log pseudolikelihood	-45930.2		
Person-observations	45743		

Reported estimates are the coefficients from a bivariate probit model. Observations are pooled over waves 8-18, with standard errors adjusted for clustering within individuals. Estimates are weighted to account for survey design and non-response and for the endpoint adjustment. * significant at 10%; ** significant at 5%.

4.2 Is there a Performance Pay Premium?

Before looking at the growth in wage dispersion in Britain and the role PP may have played we run log hourly wage regressions to establish whether there is a PP premium at the mean and, if so, how much of it can be explained by the selection of workers into jobs. The results are presented in Table 4. The first column shows the raw wage gap between those in PP jobs and those in fixed wage jobs. The second column is the regression-adjusted gap. The third column introduces person fixed effects identified using workers who switch between PP and FP jobs. The top half of the table presents results for the broad PP measure, while the bottom half focuses on the narrow PP measure.

Among men, the raw differential is about 25% ($= \exp(0.221)-1$), but falls by around half when conditioning on observable differences between PP and FP employees, and to around 4% controlling for fixed unobservable differences across employees. The pattern is similar

whether one uses the broad or narrow measure of PP. The fact that the premium falls markedly when adjusting for person fixed effects is a clear indication that there is positive selection into PP jobs among men.

Table 4 - The PP Wage Premium, 1998-2008

	Raw	Adjusted (OLS)	Adjusted (FE)
Broad PP			
Men	0.221** (0.008)	0.113** (0.006)	0.040** (0.006)
Women	0.134** (0.008)	0.086** (0.006)	0.066** (0.007)
Women (FT)	0.120** (0.010)	0.095** (0.007)	0.060** (0.008)
Narrow PP			
Men	0.228** (0.009)	0.100** (0.006)	0.044** (0.007)
Women	0.225** (0.009)	0.087** (0.006)	0.057** (0.008)
Women (FT)	0.194** (0.010)	0.098** (0.007)	0.047** (0.008)

Number of person-observations is 22,108 (men); 23,550 (women); 14,458 (FT women). Regression-adjusted estimates also control for quadratics in age and job tenure, and dummies for marital status, part-time work, temporary and fixed terms jobs, trade union coverage, public sector status, occupation (9 categories), industry (11 categories), establishment size (3 categories), region (13 categories), and wave (8-18). Raw estimates are weighted for survey design and non-response. Standard errors in parentheses. * significant at 10%; ** significant at 5%.

Turning to women, a similar pattern emerges: there is a sizeable raw wage differential which falls when regression-adjusted and falls still further with the introduction of person fixed effects. So, once again, there is clear evidence of positive sorting into PP jobs among women. This is the case among all women and women in full-time jobs. The premium when accounting for fixed unobservable differences across employees is around 6% in most cases, so a little higher than for men.⁵

To see how the PP premium may have changed over time, Table 5 reports estimates of the adjusted premium from the start and end of the time period, pooling observations from 1998-2000 and from 2006-08 respectively. The OLS estimates increased for all three subgroups on both PP measures, and some of the changes were substantial, for instance the broad PP premium for men increased from under 8% in 1998-2000 to 13% in 2006-08. The FE estimates also increased, albeit by smaller amounts and not for women working full time (for them the PP premium, on both measures, fell from 4% in 1998-2000 to an insignificant 2–3% in 2006-08). But overall there appears to be evidence of an increase in the returns to PP over the period. Whether this increase also leads to rising wage inequality will depend on where PP workers are in the wage distribution and also on how the decline in the prevalence of PP played out across the distribution. In the next section we turn to the net effect of all these factors.

⁵ It is notable that the regression-adjusted premium of around 10% is similar to the estimate Bryson et al. (2014) obtain having accounted for workplace fixed effects using the 2011 Workplace Employment Relations Survey.

Table 5 - The PP Wage Premium, 1998-2000 and 2006-08 compared

	1998-2000		2006-2008	
	OLS	FE	OLS	FE
Broad PP				
Men	0.077** (0.010)	0.049** (0.014)	0.129** (0.012)	0.060** (0.021)
Women	0.077** (0.010)	0.050** (0.017)	0.090** (0.011)	0.055** (0.022)
Women (FT)	0.070** (0.012)	0.042** (0.017)	0.101** (0.014)	0.022 (0.024)
Narrow PP				
Men	0.071** (0.010)	0.039** (0.015)	0.098** (0.012)	0.056** (0.024)
Women	0.070** (0.011)	0.065** (0.020)	0.098** (0.012)	0.070** (0.027)
Women (FT)	0.067** (0.012)	0.040** (0.018)	0.104** (0.014)	0.029 (0.027)

Number of person-observations in 1998-2000 is 6,598 (men); 6,867 (women); 4,323 (FT women). Number of observations in 2006-08 is 5,416 (men); 5,972 (women); 3,611 (FT women). Regression-adjusted estimates also control for quadratics in age and job tenure, and dummies for marital status, part-time work, temporary and fixed terms jobs, trade union coverage, public sector status, occupation (9 categories), industry (11 categories), establishment size (3 categories), region (13 categories), and wave. Standard errors in parentheses. * significant at 10%; ** significant at 5%.

4.3 Does Performance Pay Affect Wage Dispersion?

In this section we look at changes in log hourly wage dispersion between 1998-2008 in BHPS for men, women, and full-time women. First we graph dispersion in both tails of the wage distribution relative to the median. To check whether the BHPS results are consistent with other sources, we compare them to those elsewhere in the literature which tend to rely on the Annual Survey of Hours and Earnings (ASHE) and its predecessor the New Earnings Survey (NES), in particular Lindley and Machin (L&M) (2013). To ensure we have sufficient

sample sizes we use two-year moving averages.⁶ Then we present descriptive information on the mean and standard deviation in log hourly earnings for PP jobs and those in FP jobs. Finally, we compare the actual wage distribution with a counterfactual wage distribution to recover the effect of PP on wages in different parts of the wage distribution. We will explain the methodology behind this below.

Figure 2a - Dispersion in Men's Log Hourly Wages

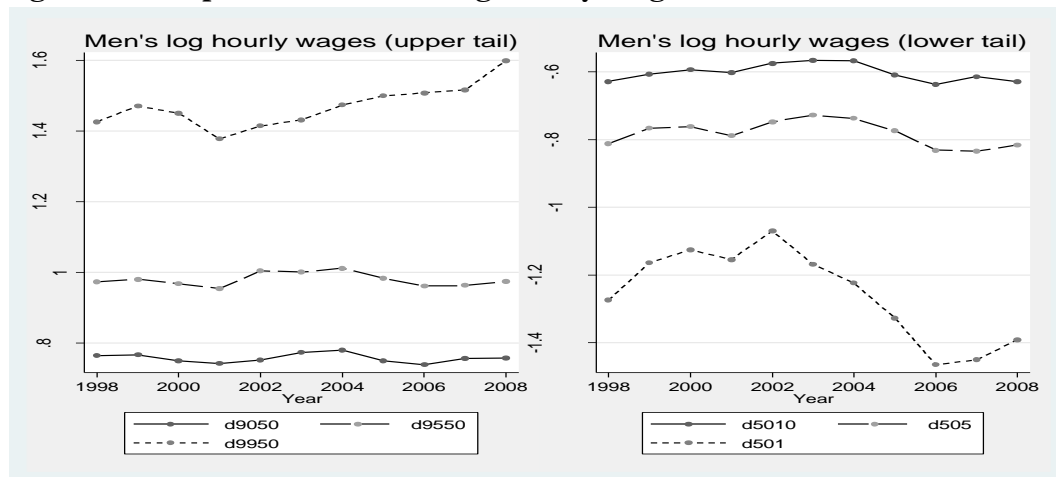
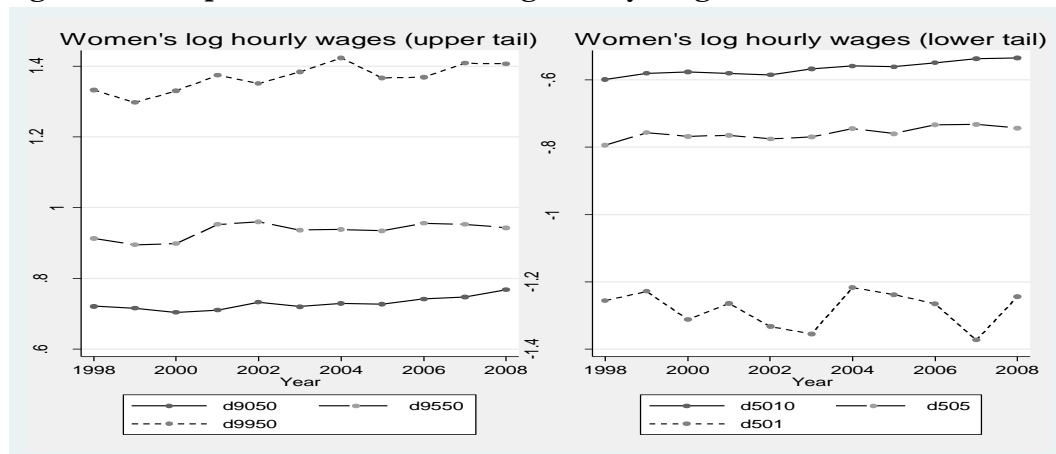
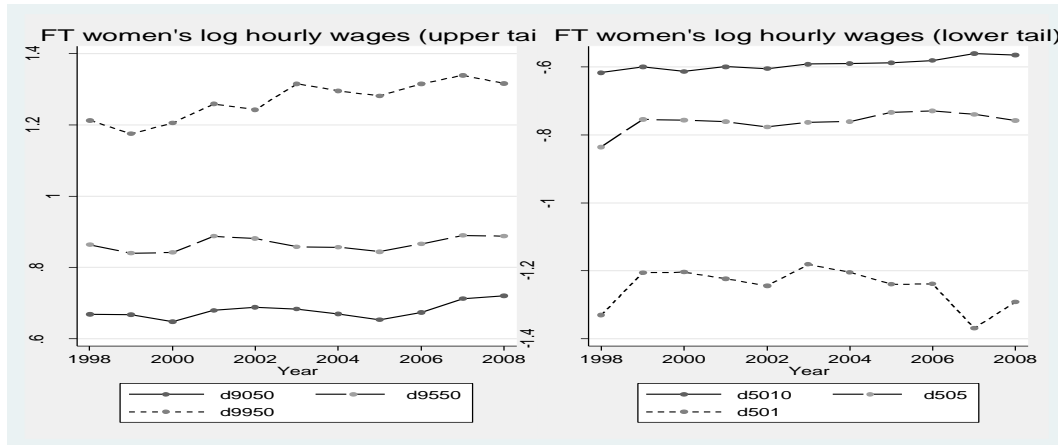


Figure 2b - Dispersion in Women's Log Hourly Wages



⁶ This means that, on average, we have around 35-40 unweighted observations in the 1% tails (25 observations for full-time women).

Figure 2c - Dispersion in Full-time Women's Log Hourly Wages



Figures 2a-2c show the log hourly wage distributions for men, women and full-time women respectively over the period 1998-2008. For men we find increasing dispersion at the very top of the distribution (the 99-50 differential), but little change further down (the 95-50 and 90-50 differentials).⁷ At the bottom we see no real change except in the 50-1 differential which fell until 2001-2 then increased sharply. In their analysis L&M find that the 90-50 differential increased over the period while the 50-10 differential reduced slightly (they do not consider the extreme tails). For women, we find increasing dispersion at the top and reducing dispersion at the bottom over 1998-2008, which is similar to the trends reported by L&M in the 90-50 and 50-10 differentials.

We therefore see some evidence of a growth in wage dispersion over the period, as do L&M, though for men there are some differences as to where precisely in the distribution this widening occurred. These differences could relate to sample differences, such as the incomplete ASHE coverage of low paid workers or the lack of coverage of new immigrants

in BHPS, or the fact that very high and very low earners are more difficult to reach with household surveys (Bollinger et al., 2014).

How is PP related to wage dispersion over the period? Table 6 shows the mean and variance of log hourly wages in the PP and FP sectors for the pooled years. It is apparent that mean wages are higher in the PP sector, in keeping with the wage premium analysis above, but there is no evidence to suggest that the variance in wages is greater in the PP sector than the FP sector, whether one is looking at men, women or full-time women. In fact, the variance for women is slightly lower in PP than FP jobs (especially by the narrow measure). One possible reason for this is that there is greater homogeneity among PP employees than there is among FP employees in terms of traits that affect their earnings. In spite of this finding, PP has the potential to affect the earnings distribution owing to the fact that employees receive a PP wage premium and, as also seen in Table 3, tend to lie higher in the earnings distribution even conditioning on other job characteristics .

Table 6 – level and dispersion of log hourly wages by job type, 1998-2008

	Fixed pay	Broad PP	Narrow PP
Mean			
Men	2.240	2.465	2.536
Women	2.089	2.225	2.325
Women (FT)	2.195	2.317	2.398
Variance			
Men	0.295	0.309	0.304
Women	0.299	0.282	0.269
Women (FT)	0.273	0.261	0.246

⁷ Similarly there is no evidence of increasing dispersion towards the middle of the distribution (not reported).

Estimates are pooled over waves 8-18 and weighted for survey design and non-response.

Our estimates of the relationship between PP and the wage distribution are based on a reweighting estimator originally deployed by DiNardo, Fortin and Lemieux (1996) and then applied in a modified form by LMP (2009). The method constructs a counterfactual wage distribution which proxies the wage distribution that would obtain in the absence of PP in the economy. This is achieved by reweighting those sample members who are not in receipt of PP such that their observable characteristics closely resemble the overall population of workers. This in turn is achieved by running a probit estimate for the probability of being in a PP job and then using the predicted probabilities to reweight the FP employees in such a way as to give additional weight to those with high estimated probabilities of being in a PP job (because these employees are underrepresented in the FP sample). One can then recover the "effect" of PP at different parts of the wage distribution by comparing the actual distribution of wages among all workers to the counterfactual distribution observed among the reweighted set of employees not in PP jobs.

Table 7 summarises the results of the counterfactual reweighting exercise for men. We consider two points in time, namely early in the period we study (1998-2000) and then again at the end of the period (2006-2008). For each time point, we report various summary measures of the actual wage distribution (columns 1 and 4) and the change in the dispersion between the two time points (column 7). The summary measures in column 7 all increased between 1998-2000 and 2006-08 (with the exception of a very small fall in the 90-50 percentile gap), although the increases are generally quite small (consistent with the graphical

analysis of the actual wage dispersion in Figure 2a), the largest being a 0.05 log point increase in the 50-5 percentile gap (column 7).⁸

In columns 2 and 5 we present the same measures of the respective counterfactual distributions. The differences between the two sets of estimates give the PP job effect at each time point (columns 3 and 6). Column 8 shows how the PP effect changed over the period. Finally, we calculate the percentage of the change in each measure of wage dispersion that can be attributed to PP jobs (column 9): this is simply the ratio of the change in the PP effect to the change in the actual wage dispersion. The table includes results for both the broad PP measure (top panel) and the narrow measure (bottom panel).⁹

PP jobs, broadly defined, account for two-thirds (64.3%) of the growth in the variance of log hourly wages over the period (row 1). The widening effect of broad PP is apparent throughout the wage distribution, but it is a little more pronounced in the upper half of the distribution, as is apparent if one compares the magnitude of the numbers in column 8. For instance, broad PP contributed to greater wage variance as measured by the 95-50 gap in both 1998-2000 (0.11 log points) and 2006-08 (0.15 log points). In contrast, broad PP actually closed the 50-5 gap in 1998-2000 (by 0.02 log points), only to see a wage dispersing effect emerge in 2006-08 (of 0.02 log points).

⁸ We do not report measures involving the 1st and 99th percentiles because the estimates were too noisy to be included. Noise in the percentile estimates is worsened by the fact that most estimates in the table involve comparing distributions (actual vs counterfactual and/or changes over time), and so the figures are differences not levels (unlike in Figure 2). Following LMP (2009) we smoothed the intermediate percentiles using a moving average of +/-2 percentiles (the unsmoothed estimates show a very similar pattern).

⁹ The measures of actual dispersion are the same in each panel but are replicated for ease of comparison with the counterfactual measures.

The picture looks very different in relation to narrow PP in the bottom panel of the table. Although narrow PP is associated with higher wage dispersion in both periods, the size of this disequalising effect fell over the period, especially at the top (so as to close the 95-50 gap by 0.03 log points). The only part of the distribution where the disequalising effect grew was near the bottom, and this effect is quantitatively small (widening the 50-5 gap by 0.01 log points).

These divergent results for broad and narrow PP highlight the value in distinguishing between various types of PP when evaluating its impact on wages and wage dispersion. The results here confirm the importance of bonus payments in particular in understanding PP effects on growing wage dispersion among men. One possible reason why bonuses are particularly important in affecting wage dispersion is that bonuses are particularly prevalent at the top of the wage distribution and, as other research using the BHPS has shown, among men bonuses tend to substitute for fixed pay at the bottom of the wage distribution but this substitution effect is not apparent at the top of the distribution (Green and Heywood, 2012).¹⁰

¹⁰ Close to two-fifths of male employees received bonuses during this period. BHPS also asks "What was the total amount of bonus you received over the last twelve months?" For those receiving a bonus they were the equivalent of around about 3.5% of base pay in the late 1990s, rising to 4.5-5% towards the end of our period of investigation.

Table 7 – Effect of PP jobs on log wage distribution (men), 1998-2000 and 2006-2008

	1998-2000			2006-08			1998-2000 to 2006-08		
	Actual dispersion (1)	Dispersion without PP jobs (2)	PP job effect (2)-(1) (3)	Actual dispersion (4)	Dispersion without PP jobs (5)	PP job effect (5)-(4) (6)	Change in actual dispersion (4)-(1) (7)	Change in PP job effect (6)-(3) (8)	Percentage of dispersion change due to PP jobs (8)/(7) (9)
<i>Broad PP measure</i>									
Variance	0.304	0.281	0.023	0.344	0.295	0.048	0.039	0.025	64.3
Percentile gaps									
90-10	1.373	1.308	0.065	1.389	1.306	0.083	0.016	0.018	110.6
95-50	0.981	0.874	0.106	1.006	0.853	0.153	0.025	0.047	186.8
90-50	0.761	0.688	0.074	0.760	0.681	0.080	-0.001	0.006	-724.4
50-10	0.612	0.620	-0.008	0.629	0.625	0.004	0.017	0.012	69.9
50-5	0.790	0.812	-0.022	0.844	0.826	0.018	0.054	0.040	73.9
<i>Narrow PP measure</i>									
Variance	0.304	0.287	0.017	0.344	0.334	0.010	0.039	-0.007	-17.6
Percentile gaps									
90-10	1.373	1.324	0.048	1.389	1.365	0.025	0.016	-0.024	-145.8
95-50	0.981	0.927	0.054	1.006	0.980	0.025	0.025	-0.029	-115.4
90-50	0.761	0.721	0.040	0.760	0.735	0.026	-0.001	-0.015	1777.0
50-10	0.612	0.604	0.008	0.629	0.630	-0.001	0.017	-0.009	-52.1
50-5	0.790	0.785	0.005	0.844	0.826	0.018	0.054	0.014	25.4

While Table 7 presents estimates of the size of PP effects on wage dispersion among men at specific points of the wage distribution, we can also illustrate the distributional effects using graphs. Figures 3a and 3b present the effects of broad PP over the wage distribution for men. The solid line in Figure 3a represents the difference broad PP made to log hourly wage dispersion in the period 2006-08 by comparing the actual log hourly wages of all male employees - who are a mixture of PP and FP workers - with counterfactual wages based on a scenario in which nobody receives PP (the corresponding summary measures are in the top panel of Table 7 in column 3). The counterfactual gap is fairly flat in the bottom half of the wage distribution, but then it begins to rise such that the log wage differential is around .15 log points towards the top of the wage distribution.

The dotted line presents the same information but for the period 2006-2008 (corresponding to column 6 in the top panel of Table 7). The effect of PP is more pronounced in the later period, rising much more steeply in the top half of the wage distribution. Consistent with the summary measures reported above, the graphs indicate that PP has a disequalising effect on wages which increased in the later period.

Figure 3a - Effect of PP jobs (broad measure) on wage distribution in 1998-2000 and 2006-2008, Men

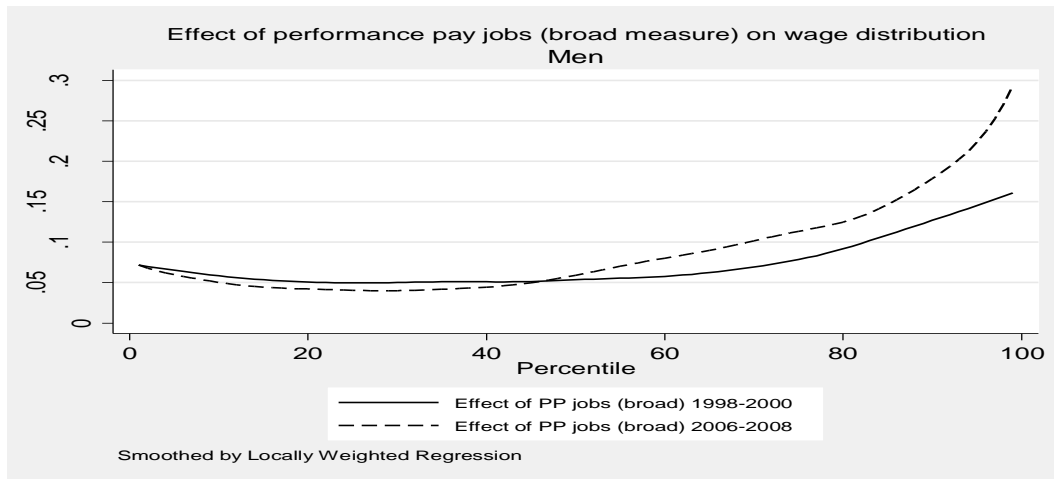


Figure 3b - Effect of PP jobs (broad measure) on change in wage distribution between 1998-2000 and 2006-2008, Men

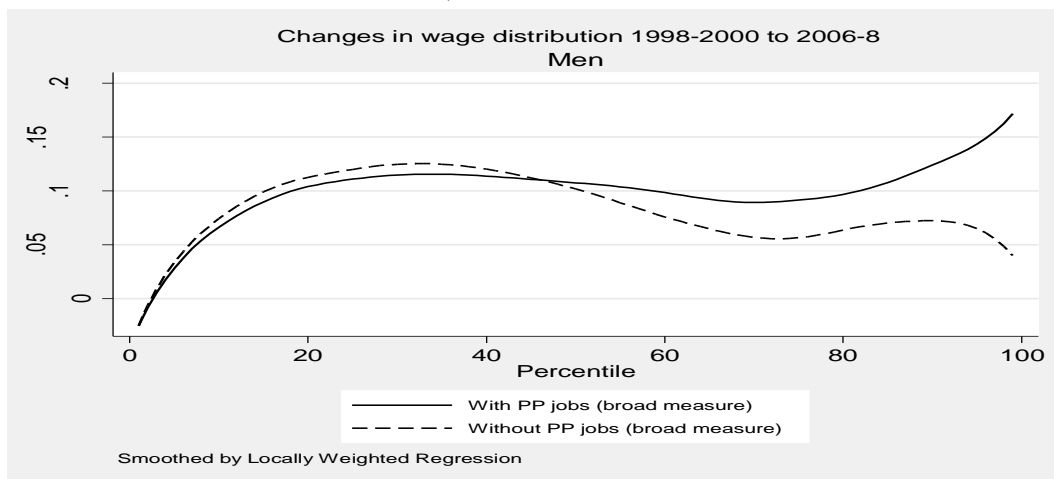


Figure 3b shows the role played by broad PP in the *change* in the male wage distribution between the early and late periods. The solid line shows how the male wage distribution actually changed (corresponding to column 7 in the top panel in Table 7), while the dotted line shows how the male wage distribution would have changed in the absence of PP (that is, the difference in the counterfactual scenarios for 1998-2000 and 2006-08). Thus the gap between the two lines gives the effect of PP on changes in the distribution (corresponding to column 8 in the upper panel of Table 7). PP makes little difference to the change in the wage

distribution in the lower half of the wage distribution: wage dispersion grew in the lowest quartile of the distribution, and would have done in a similar fashion in the absence of PP. We noted above that PP widened the 50-5 gap (relative to the counterfactual without PP jobs) but it is clear from the graph that most of the change was due to a small rise in the median (the solid line is higher than the dotted one) and not a fall in the 5th percentile. This illustrates how the graph can provide a more complete picture than just comparing two points alone.

In contrast, the graph confirms the figures in Table 7 showing broad PP contributed to rising earnings dispersion in the top half of the wage distribution: in the absence of PP the wage growth in the upper part of the distribution would have been about half of what it actually was (an increase of about .10 log points compared with about .05 log points at the 80th percentile, for example). In fact, what is striking from the graph is that wage dispersion would have actually fallen between roughly the median and the 80th percentile without PP jobs. Higher up (between about the 80th and 90th percentiles), dispersion would have increased to a small extent, but actual dispersion, reflecting the effect of PP jobs, increased much more.

As indicated in Table 7, the picture looks rather different for men if we focus on the narrow measure of PP. PP measured in this way does result in a wider wage dispersion than would be the case in its absence. Although this was the case both at the beginning and the end of our period of investigation, the effect was attenuated in the second period (Figure 4a).

Consequently, the effect of narrow PP jobs on changes in the wage distribution over the period was actually to reduce that dispersion, though not by very much (Figure 4b).¹¹

Figure 4a - Effect of PP jobs (narrow measure) on wage distribution in 1998-2000 and 2006-2008, Men

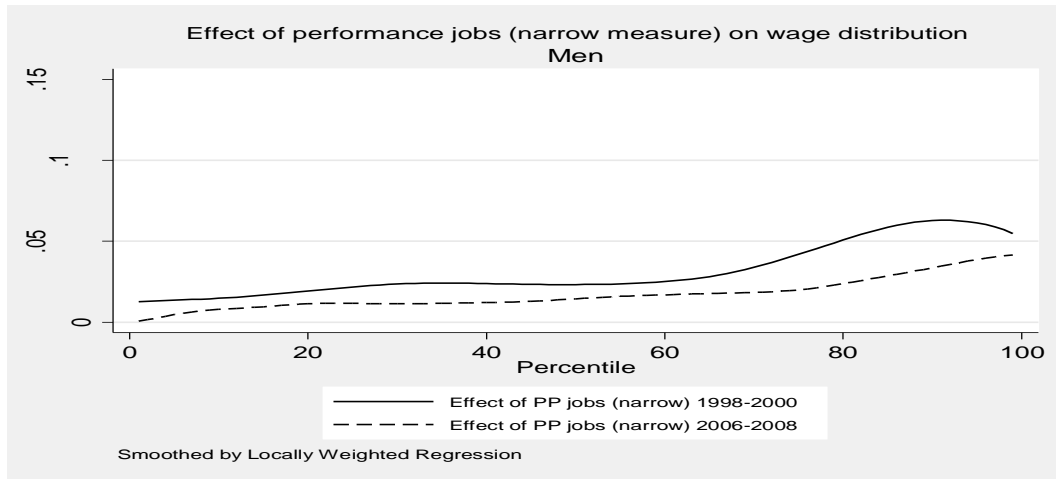
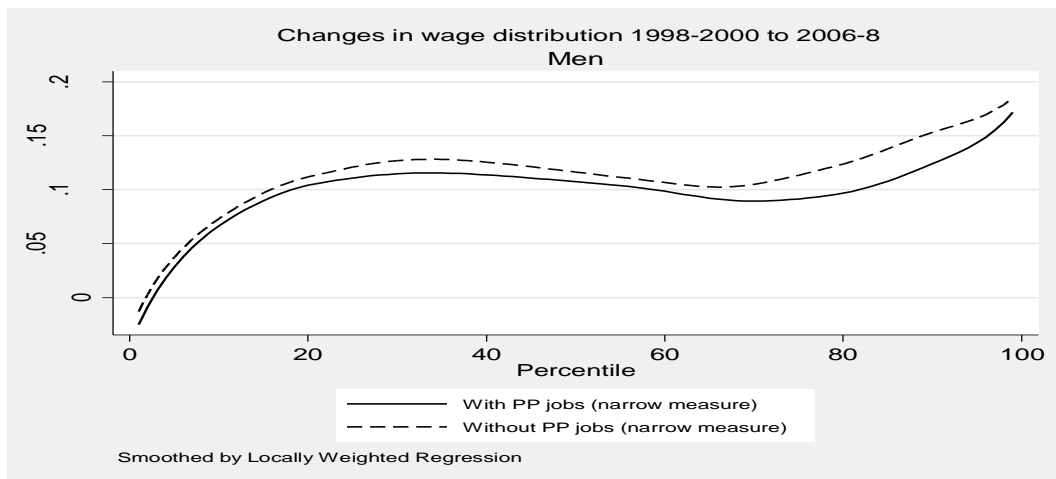


Figure 4b - Effect of PP jobs (narrow measure) on change in wage distribution between 1998-2000 and 2006-2008, Men



¹¹ As noted earlier, close to two-fifths of male employees received bonuses during this period. BHPS also asks "What was the total amount of bonus you received over the last twelve months?" For those receiving a bonus they were the equivalent of around about 3.5% of base pay in the late 1990s, rising to 4.5-5% towards the end of our period of investigation.

Table 8 – Effect of PP jobs on log wage distribution (women), 1998-2000 and 2006-2008

	1998-2000			2006-08			1998-2000 to 2006-08		
	Actual dispersion (1)	Dispersion without PP jobs (2)	PP job effect (2)-(1) (3)	Actual dispersion (4)	Dispersion without PP jobs (5)	PP job effect (5)-(4) (6)	Change in actual dispersion (4)-(1) (7)	Change in PP job effect (6)-(3) (8)	Percentage of dispersion change due to PP jobs (8)/(7) (9)
<i>Broad PP measure</i>									
Variance	0.280	0.296	-0.016	0.291	0.294	-0.004	0.011	0.013	118.7
Percentile gaps									
90-10	1.293	1.316	-0.023	1.306	1.328	-0.021	0.013	0.001	10.5
95-50	0.917	0.918	-0.001	0.959	0.926	0.034	0.042	0.034	80.8
90-50	0.712	0.711	0.001	0.764	0.755	0.009	0.052	0.008	14.9
50-10	0.582	0.606	-0.024	0.542	0.573	-0.030	-0.039	-0.006	16.3
50-5	0.772	0.806	-0.034	0.752	0.813	-0.061	-0.021	-0.028	133.8
<i>Narrow PP measure</i>									
Variance	0.280	0.285	-0.005	0.291	0.284	0.007	0.011	0.012	113.5
Percentile gaps									
90-10	1.293	1.305	-0.012	1.306	1.291	0.016	0.013	0.028	211.5
95-50	0.917	0.907	0.010	0.959	0.933	0.026	0.042	0.016	38.6
90-50	0.712	0.710	0.002	0.764	0.743	0.021	0.052	0.019	35.8
50-10	0.582	0.595	-0.014	0.542	0.547	-0.005	-0.039	0.009	-22.6
50-5	0.772	0.789	-0.017	0.752	0.763	-0.011	-0.021	0.006	-28.6

Now we turn to wage dispersion among women. It is apparent from column 7 in Table 8 that overall wage dispersion among women grew over the period, but only very marginally (variance increased by 0.011). This is partly because trends went in opposite directions in the top and bottom halves of the wage distribution: wages became more compressed in the bottom half of the wage distribution, whereas they became more dispersed in the top half of the distribution. How did PP affect this distribution?

If we begin with the broad PP measure, it is apparent that it tended to compress women's earnings at the bottom of the distribution in both periods (column 3 for 1998-2000 and column 6 for 2006-2008). This compressing effect became more apparent in the second period. There was little effect at the top of the distribution in 1998-2000 (column 3) but some widening in 2006-2008 (column 8).

The overall effects of these countervailing effects of the broad measure of PP on women's wages is best illustrated graphically. Figure 5a indicates that the broad measure of PP resulted in a wage distribution for women that was more U-shaped than it would have been in its absence. However, the U was flatter further up the wage distribution in the second period relative to the first. This is why PP contributed to a growth in wage dispersion in the top half of the wage distribution compared with a counterfactual world without PP (compare the solid line with the dotted line in Figure 5b). This pattern of results is similar for women when using the narrow measure of PP jobs (Figures 6a and 6b).

Figure 5a - Effect of PP jobs (broad measure) on wage distribution in 1998-2000 and 2006-2008, All Women

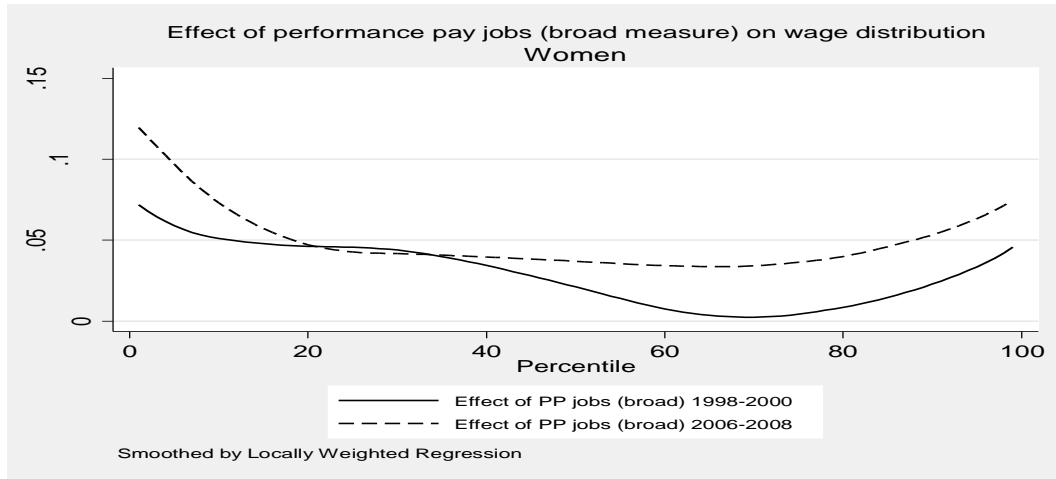


Figure 5b - Effect of PP jobs (broad measure) on change in wage distribution between 1998-2000 and 2006-2008, All Women

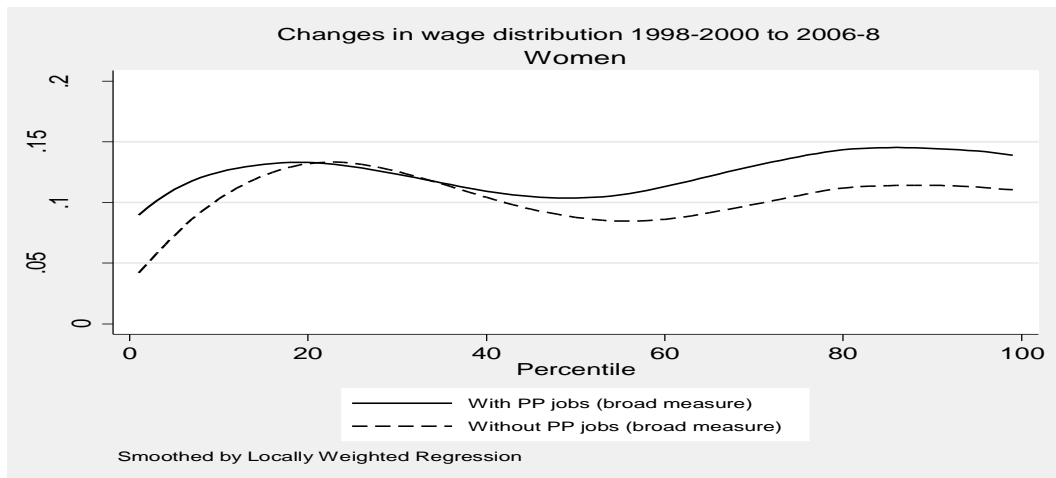


Figure 6a - Effect of PP jobs (narrow measure) on wage distribution in 1998-2000 and 2006-2008, All Women

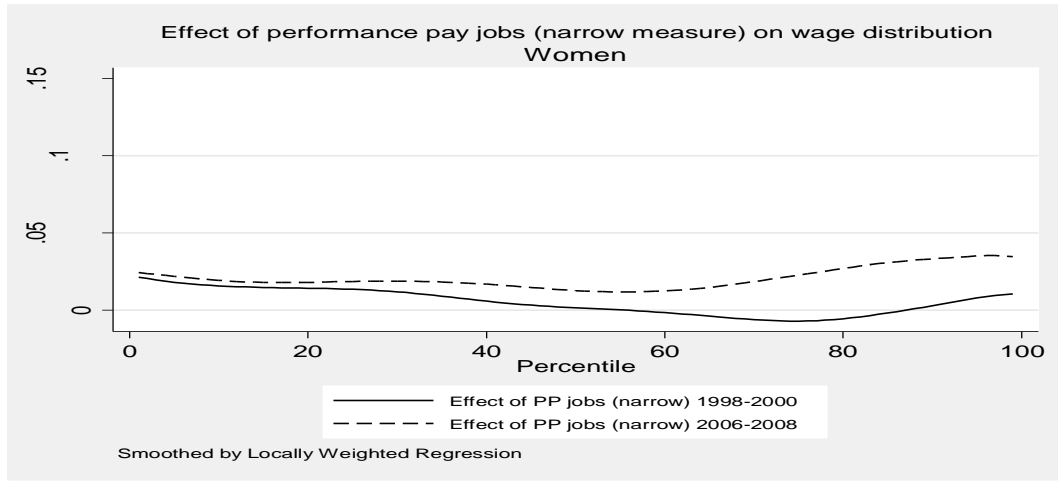


Figure 6b - Effect of PP jobs (narrow measure) on change in wage distribution between 1998-2000 and 2006-2008, All Women

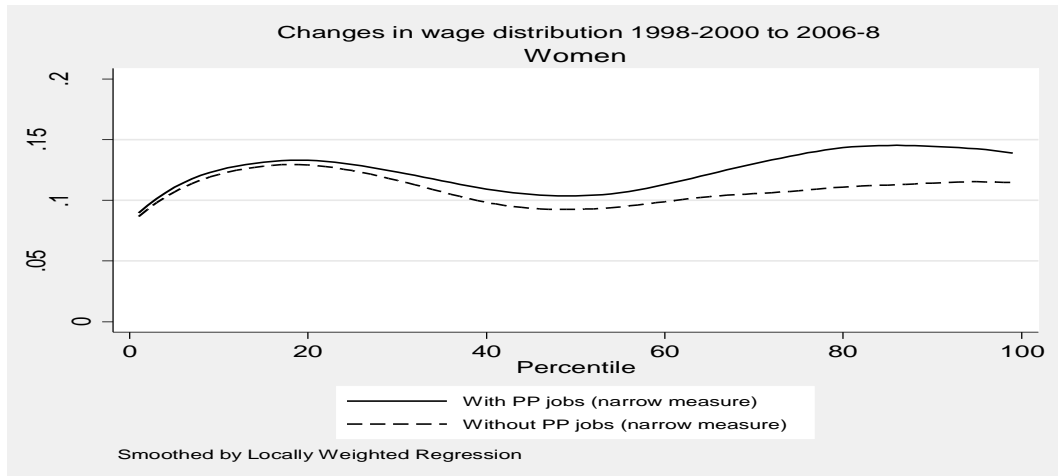


Table 9 – Effect of PP jobs on log wage distribution (Full-time women), 1998-2000 and 2006-2008

	1998-2000			2006-08			1998-2000 to 2006-08		
	Actual dispersion (1)	Dispersion without PP jobs (2)	PP job effect (2)-(1) (3)	Actual dispersion (4)	Dispersion without PP jobs (5)	PP job effect (5)-(4) (6)	Change in actual dispersion (4)-(1) (7)	Change in PP job effect (6)-(3) (8)	Percentage of dispersion change due to PP jobs (8)/(7) (9)
<i>Broad PP measure</i>									
Variance	0.252	0.257	-0.005	0.274	0.286	-0.012	0.022	-0.007	-31.2
Percentile gaps									
90-10	1.263	1.282	-0.019	1.277	1.290	-0.013	0.014	0.006	40.0
95-50	0.849	0.803	0.046	0.907	0.861	0.046	0.058	0.001	1.0
90-50	0.660	0.630	0.030	0.713	0.677	0.036	0.053	0.006	11.9
50-10	0.603	0.652	-0.049	0.564	0.613	-0.049	-0.040	-0.001	2.1
50-5	0.778	0.832	-0.055	0.766	0.873	-0.107	-0.011	-0.052	466.0
<i>Narrow PP measure</i>									
Variance	0.252	0.255	-0.002	0.274	0.273	0.001	0.022	0.003	15.3
Percentile gaps									
90-10	1.263	1.266	-0.003	1.277	1.257	0.019	0.014	0.022	160.8
95-50	0.849	0.825	0.024	0.907	0.873	0.034	0.058	0.010	17.2
90-50	0.660	0.641	0.019	0.713	0.669	0.044	0.053	0.025	47.7
50-10	0.603	0.625	-0.021	0.564	0.588	-0.025	-0.040	-0.003	8.1
50-5	0.778	0.801	-0.024	0.766	0.805	-0.039	-0.011	-0.015	132.7

Finally we turn to women in full-time employment. Their wages became more dispersed over the period, the effect being more pronounced than for all women (variance rose by 0.02, Table 9, column 7, compared with 0.01 for all women, Table 8, column 7). However, as in the case of all women, there was compression in earnings at the bottom of the distribution and growing earnings dispersion at the top of the distribution. The growth in dispersion in the top half of the distribution is actually quite substantial: the 95-50 ratio grows by 0.06 log points.

Turning to the effects of PP on the dispersion of full-time women's earnings and focusing first on the broad PP measure it is apparent that PP is associated with greater wage dispersion at the top of the distribution but lower dispersion at the bottom of the distribution. This is the case in both 1998-2000 and 2006-2008 (upper panel Table 9, columns 3 and 6). These effects change very little over the whole period, with the exception of the compressing effect of broad PP on the 50-5 differential, which almost doubles.

Looking at the whole distribution graphically using the broad PP measure the PP effect on full-time women's earnings relative to a counterfactual world without PP is highest at the top and bottom of the earnings distribution, forming the U-shape referred to above for all women. The size of this effect is larger in the second period (2006-2008) relative to the early period (1998-2000) but it is similar across all the distribution, except at the very top (Figure 7a). For this reason PP (broadly defined) resulted in higher earnings among full-time working women, but it had little effect on changing inequality (which decreased below the

10th percentile and increased above the median), except perhaps to mitigate the increase at the very top (Figure 7b).

Figure 7a - Effect of PP jobs (broad measure) on wage distribution in 1998-2000 and 2006-2008, Full-time women

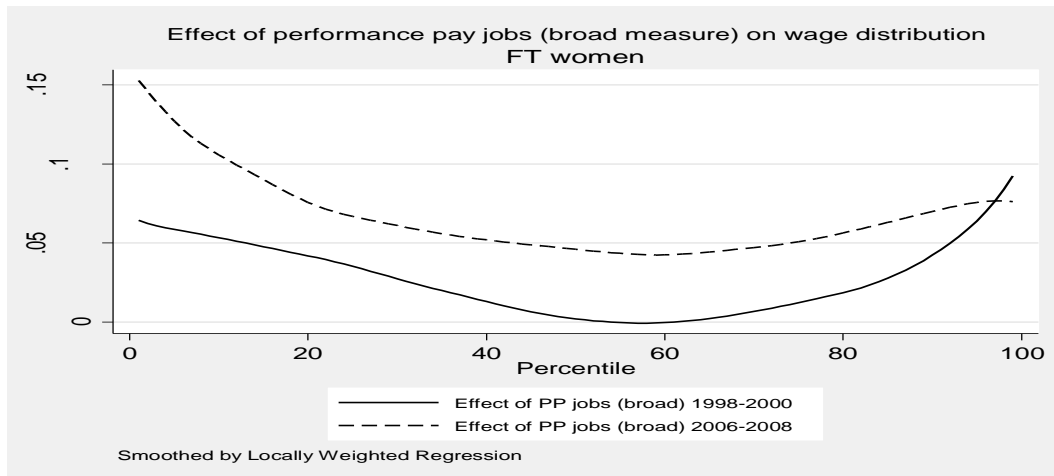
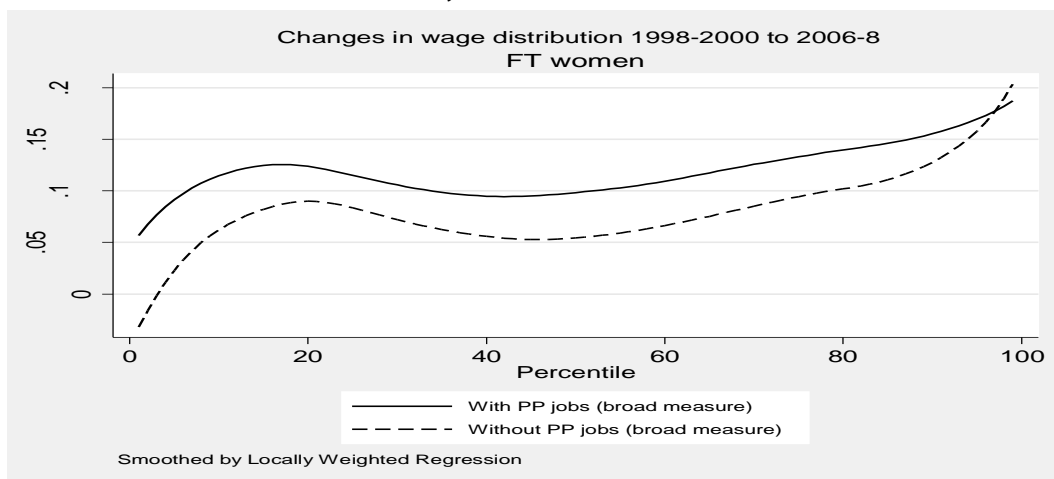


Figure 7b - Effect of PP jobs (broad measure) on change in wage distribution between 1998-2000 and 2006-2008, Full-time women



If we turn to the narrow PP measure and consider its effects on the log hourly earnings of women working full-time this is a shallow U-shape, in both periods, but the size of the effect is greater in 2006-08 (Figure 8a) such that PP increases wage dispersion over the period, as

indicated by the rising solid line in Figure 8b relative to the dotted line for the counterfactual "no PP" world, once one moves beyond the lowest quartile of the earnings distribution.

Figure 8a - Effect of PP jobs (narrow measure) on wage distribution in 1998-2000 and 2006-2008, Full-time Women

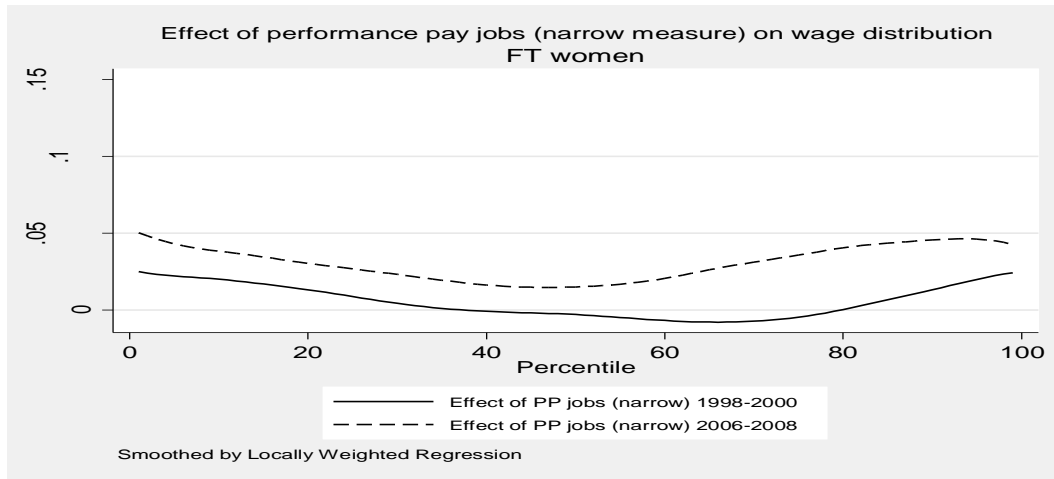
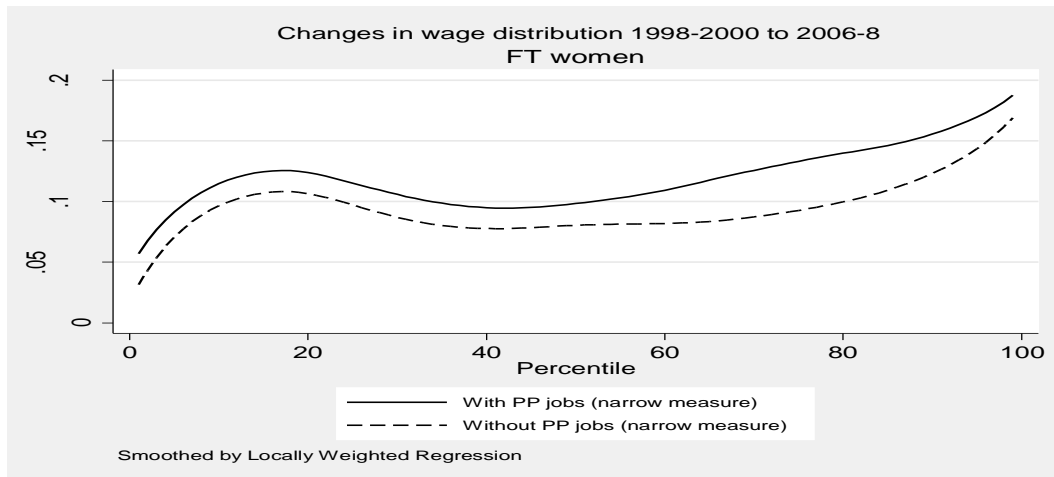


Figure 8b - Effect of PP jobs (narrow measure) on change in wage distribution between 1998-2000 and 2006-2008, Full-time women



5. Conclusions

There has been much speculation about the various causes of growing wage dispersion in Britain, the United States and elsewhere. The seminal paper by LMP (2009) showed PP contributed significantly to the growth in earnings dispersion in the United States through to

the early 1990s. Using data from the British Household Panel Survey (BHPS) we adopt a similar estimation approach to LMP but applied to Britain during the decade of economic growth that ended abruptly with the recession of 2008. In contrast to LMP, we find that rather than increasing, the prevalence of PP declined (or at most stayed flat) over the decade to 2008. This applies to both broad and narrow measures of PP (although there is some evidence that bonus payments increased), and the trends appear comparable with more recent declines in the US identified by Gittleman and Pierce (2013).

We confirm others' findings that wage inequality grew overall during the decade to 2008, largely due to growing earnings dispersion in the top half of the wage distribution, but there was also some reduction of inequality among women in the bottom of the distribution. The contribution of PP to these changes depends on how the incidence and returns to PP changed and where workers sit in the wage distribution. While the incidence of PP fell, there was still a substantial wage return to PP and indeed it appears to have increased over time. Comparing the actual wage distribution with a counterfactual world without PP, we find the net effect of these changes to be that PP is associated with greater wage dispersion towards the top and, particularly for broad PP, that this disequalising effect increased over the period, possibly because of increased bonuses. This was accompanied by some counterveiling effects at the bottom, in particular broad PP is associated with more compressed wages in the lower half of the distribution for women.

Overall PP contributed to earnings dispersion for both men and women in the upper half of the distribution although most of the growth was attributable to a particular type of PP namely bonuses. PP also explains some of the reduction in inequality among women in the

lower half of the wage distribution. However, the size of these PP effects is relatively modest, perhaps in part because, in contrast to LMP, there is no indication that PP jobs became increasingly prevalent, among either men or women, in the decade prior to recession.

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